These problems are meant to test your understanding of the concepts in Module 1. They are *not* to be handed in. Some of these have been modified (or in some cases taken directly) from questions in the *Additional Resources* listed in the course syllabus, and no claims of originality are made.

1. Suppose  $\mathcal{X} = \{1, 2, 3, 4\}$  and  $\Theta = \{a, b\}$ . Two mass functions on  $\mathcal{X}$  – one for each value of  $\theta \in \Theta$  – are specified in the following table:

	x = 1	x = 2	x = 3	x = 4
$f_a(x)$	1/2	1/6	1/6	1/6
$f_b(x)$	1/4	1/4	1/4	1/4

Suppose  $X \sim f_{\theta}$ , and let  $T(x) = \mathbb{1}_{x \in \{2,3,4\}}$ .

Going from the definition, show that T(X) is a sufficient statistic for  $\theta$ . You can do this by working out a table of  $\mathbb{P}_a(X = x \mid T(X) = t)$  for each  $x \in \mathcal{X}$  and  $t \in \{0, 1\}$ , and then doing the same for  $\mathbb{P}_b(X = x \mid T(X) = t)$ .

- 2. Let  $X \sim \mathcal{N}(0, \sigma^2)$ , where  $\sigma^2 > 0$ . Prove that T(X) = |X| is sufficient for  $\sigma^2$ .
- 3. Suppose  $X_1, X_2, \ldots, X_n \stackrel{iid}{\sim} f_{\theta}$ , where  $\theta \in \Theta$  is unknown. Show that  $T(\mathbf{X}) = (X_{(1)}, X_{(2)}, \ldots, X_{(n)})$  is sufficient for  $\theta$ .
- 4. Let  $X_1, X_2, \ldots, X_n \stackrel{iid}{\sim} \text{Unif}[0, \theta]$ , where  $\theta > 0$ . Find a minimal sufficient statistic for  $\theta$ .
- 5. Let  $X_1, X_2, \ldots, X_n \stackrel{iid}{\sim} \text{Unif}[\theta, \theta + 1]$ , where  $\theta \in \mathbb{R}$ . Find a minimal sufficient statistic for  $\theta$ .
- 6. Let  $X_1, X_2, \ldots, X_n \stackrel{iid}{\sim} \text{Unif}[\theta_1, \theta_2]$ , where  $\theta_1, \theta_2 \in \mathbb{R}$  and  $\theta_1 < \theta_2$ . Find a minimal sufficient statistic for  $\theta = (\theta_1, \theta_2)$ .
- 7. Let  $X_1, X_2, \ldots, X_n \stackrel{iid}{\sim} \text{Geometric}(\theta)$ , where  $\theta \in (0, 1)$ . Find a complete sufficient statistic for  $\theta$ .
- 8. Let  $X_1, X_2, \ldots, X_n$  be a random sample from a continuous distribution with density

$$f_{\theta}(x) = \theta x^{\theta - 1}, \quad 0 < x < 1, \quad \theta > 0$$

Show that  $T(\mathbf{X}) = \prod_{i=1}^{n} X_i$  is a complete sufficient statistic for  $\theta$ .

9. This will give you some practice dealing with multi-parameter exponential families (which is a hint!). Let  $X_1, X_2, \ldots, X_n$  be a random sample from an inverse Gaussian distribution, which has density

$$f_{\mu,\lambda}(x) = \sqrt{\frac{\lambda}{2\pi x^3}} \exp\left(-\frac{\lambda(x-\mu)^2}{2\mu^2 x}\right), \quad x \in \mathbb{R}, \quad \lambda > 0, \quad \mu \in \mathbb{R}.$$

Show that

$$T(\mathbf{X}) = \left(\bar{X}_n, \sum_{i=1}^n X_i \cdot \sum_{i=1}^n \frac{1}{X_i}\right)$$

is a complete sufficient statistic for  $\theta := (\lambda, \mu)$ .

10. Let  $X_1, X_2, \ldots, X_n$  be a random sample from a Beta $(\alpha, k\alpha)$  distribution with  $k \in \mathbb{N}$  known, which has density

$$f_{\alpha}(x) = \frac{\Gamma((k+1)\alpha)}{\Gamma(\alpha)\Gamma(k\alpha)} x^{\alpha-1} (1-x)^{k\alpha-1}, \quad x \in (0,1), \quad \alpha > 0.$$

Find a complete sufficient statistic for  $\alpha$ .

11. Let  $X_1, X_2, \ldots, X_n$  be a random sample from a Pareto( $\theta$ ) distribution, which has density

$$f_{\theta}(x) = \frac{\theta x_0^{\theta}}{x^{\theta+1}}, \quad x \ge x_0, \quad \theta > 1.$$

Here  $x_0 > 0$  is known. Find a complete sufficient statistic for  $\theta$ .

12. Using our notation, the Evans/Rosenthal textbook initially defines a sufficient statistic like this:

A function  $T(\cdot)$  defined on the sample space  $\mathcal{X}^n$  is a *sufficient statistic* for  $\theta$  if the following holds:  $T(\mathbf{x}) = T(\mathbf{y})$  implies  $f_{\theta}(\mathbf{x}) = c(\mathbf{x}, \mathbf{y}) \cdot f_{\theta}(\mathbf{y})$ , for some constant  $c(\mathbf{x}, \mathbf{y}) > 0$ .

It turns out that their definition is equivalent to ours, but proving that fact is fairly difficult.

- (a) Instead, just prove that our definition implies theirs. You can stick to the discrete case. The trick is to observe that the event  $\{\mathbf{X} = \mathbf{x}\}$  is a subset of the event  $\{T(\mathbf{X}) = T(\mathbf{x})\}$ .
- (b) If we replace the word "implies" in the textbook definition with "if and only if", what can be said about  $T(\cdot)$ ?
- 13. Prove that if a statistic  $T(\mathbf{X})$  is complete for  $\theta$  and r is a bijection, then the statistic  $r(T(\mathbf{X}))$  is also complete for  $\theta$ .
- 14. Suppose  $\mathcal{X} = \{0, 1, 2\}$  and  $\Theta = (0, \frac{1}{2})$ .

	x = 0	x = 1	x = 2
$f_{\theta}(x)$	$\theta$	$\theta^2$	$1-\theta-\theta^2$

Let  $X \sim f_{\theta}$ . Prove that T(X) = X is a complete sufficient statistic for  $\theta$ .

- 15. Let  $X_1, X_2, \ldots, X_n$  be a random sample from a scale family with parameter  $\sigma > 0$ . Prove that any function of the n-1 ratios  $X_1/X_n, \ldots, X_{n-1}/X_n$  must be ancillary for  $\sigma$ . Hint:  $X_i/\sigma \sim F(x)$ .
- 16. Let  $X_1, X_2, \ldots, X_n$  be a random sample from a location family with parameter  $\mu \in \mathbb{R}$ . Prove that any function of the n-1 differences  $X_1 X_n, X_2 X_n, \ldots, X_{n-1} X_n$  is ancillary for  $\mu$ .
- 17. Let  $X_1, X_2, \ldots, X_n \stackrel{iid}{\sim} \mathcal{N}(\mu, \sigma^2)$  with  $\mu \in \mathbb{R}$  and  $\sigma^2 > 0$ . Show that  $R(\mathbf{X}) = (X_1 \bar{X}, X_2 \bar{X}, \ldots, X_n \bar{X})$  is ancillary for  $\mu$ . Is  $R(\mathbf{X})$  independent of  $\bar{X}$ ? This example will be very important in Module 4.
- 18. We've not had any issues checking the "open set" condition of Theorem 1.8, but here's a famous example that shows you what can happen if it's not satisfied. Let  $X_1, X_2, \ldots, X_n \stackrel{iid}{\sim} \mathcal{N}(\theta, \theta^2)$ , where  $\theta > 0$ . One can show that the condition doesn't hold for this parameter space (go ahead and try if you know a bit of topology; otherwise, don't bother).
  - (a) Show that in this case, the statistic in the theorem is

$$T(\mathbf{X}) = \left(\sum_{i=1}^{n} X_i, -\frac{1}{2} \sum_{i=1}^{n} X_i^2\right)$$

- (b) Show that  $T(\mathbf{X})$  is a one-to-one function of  $(\bar{X}_n, S^2)$ .
- (c) Show that  $\mathbb{E}_{\theta}\left[\frac{n}{n+1}(\bar{X}_n)^2 S^2\right] = 0$  for all  $\theta > 0$ .
- (d) Clearly, it's not always true that  $\frac{n}{n+1}(\bar{X}_n)^2 \neq S^2$  (try it with a few small numbers if you're skeptical). Explain why this implies that  $T(\mathbf{X})$  can't be complete for  $\theta$ .
- 19. Show that the following distributions are in exponential families, assuming all parameters are unknown unless otherwise specified. For each one, identify the parameter  $\theta$  (which may well be a vector), and the functions h(x),  $g(\theta)$ ,  $T_j(x)$ , and  $\eta_j(\theta)$  for each j (if there's more than one). Also decide whether any belong to a location family, a scale family, or a location-scale family.
  - (a)

$$f_{\mu,\sigma}(x) = \frac{1}{x\sigma\sqrt{2\pi}} \exp\left(-\frac{(\log(x) - \mu)^2}{2\sigma^2}\right), \quad x > 0, \qquad \mu \in \mathbb{R}, \quad \sigma > 0.$$

(b)

$$f_{k,\lambda}(x) = \frac{\lambda^k x^{k-1} e^{-\lambda x}}{(k-1)!}, \quad x > 0, \qquad k \in \mathbb{N}, \quad \lambda > 0.$$

(c)

$$f_{\nu}(x) = \frac{2^{-\nu/2} x^{-(\nu/2+1)}}{\Gamma(\nu/2)} e^{-\frac{1}{2x}}, \quad x > 0, \qquad \nu > 0.$$

(d)

$$p_{\theta}(x) = \binom{x+r-1}{x} (1-\theta)^r \theta^x, \quad x \in \{0, 1, 2, \ldots\}, \qquad r \text{ known}, \quad \theta \in [0, 1].$$

(e)

$$p_{\mathbf{p}}(\mathbf{x}) = \binom{n}{x_1, \dots, x_k} \prod_{i=1}^k p_i^{x_i}, \quad \mathbf{x} \in \{0, 1, \dots, n\}^k \text{ s.t. } \sum_{i=1}^k x_i = n, \quad \mathbf{p} \in (0, 1)^k \text{ s.t. } \sum_{i=1}^k p_i = 1,$$

with n known.

20. Is it possible to have a (continuous) scale family  $\{f_{\sigma}(\cdot) : \sigma > 0\}$  where each  $f_{\sigma}$  is supported on the *same* bounded interval [a, b]? If so, come up with one. If not, prove it.