

UNIVERSITY OF TORONTO  
Faculty of Arts and Science

STA261H1: Probability and Statistics II  
Midterm 2  
July 30, 2024

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- Do not open this test until you are told to begin.
  - Midterm 2 is closed-book; no aids are allowed.
  - Whatever you write on the backs of these pages will not be graded, so you can use them for scratch work.
  - There are six questions (worth a total of 30 points) on the midterm, plus one bonus question (worth five additional points). Take a quick scan through the questions first and prioritize your time accordingly.
  - Show all of your work for full marks, and ensure your notation is legible, correct, and consistent with that used in the course. Use  $\vec{X}$  and  $\vec{x}$  to denote vectors.
  - If you need to use a result from lecture, either refer to it by its name (if it is a named theorem), or briefly describe it.

Good luck!

1. (5 points) Answer the following.

(a) (1 point) State the definition of a **pivotal quantity**.

(b) (1 point) State the **Neyman-Pearson lemma**.

(c) (2 point) Describe what is meant by a **goodness of fit test**, and give an example of one.

(d) (1 point) State what is meant by a **type II error**.

2. (5 points) Let  $X_1, X_2, \dots, X_n \stackrel{iid}{\sim} \text{Unif}(-\theta, \theta)$ , where  $\theta \in \mathbb{R}$ .<sup>1</sup> Suppose we want to test  $H_0 : \theta = \theta_0$  versus  $H_A : \theta \neq \theta_0$  using a test characterized by the rejection region  $R = \{\mathbf{x} \in \mathcal{X}^n : x_{(n)} > c\}$  for some  $c > 0$ .

(a) (2.5 points) Compute the power function of the test.

(b) (2.5 points) Find the  $c$  that makes this a size- $\alpha$  test.

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<sup>1</sup>Recall that when  $a < b$ , the general  $\text{Unif}(a, b)$  distribution has cdf given by  $F_{a,b}(x) = \frac{b-x}{b-a} \cdot \mathbb{1}_{x \in (a,b)} + \mathbb{1}_{x \geq b}$ . Also, if  $X_1, \dots, X_n$  are iid and continuous with cdf  $F_X$ , then  $X_{(n)}$  has cdf  $F_{X_{(n)}}(x) = (F_X(x))^n$ .

3. (5 points) Assume that  $X_1, X_2, \dots, X_n \stackrel{iid}{\sim} \mathcal{N}(\mu, \sigma^2)$ , where  $\mu \geq 1$  and  $\sigma^2$  is known.<sup>2</sup>
- (a) (2.5 points) Find the value of  $b \in \mathbb{R} \setminus \{0\}$  that makes  $(\bar{X}_n/b, \infty)$  into a lower one-sided  $(1 - \alpha)$ -confidence interval for  $\mu$ .

- (b) (2.5 points) Even though we don't know  $\mu$ , explain how we might use a visual diagnostic to assess the assumption that the data is  $\mathcal{N}(\mu, \sigma^2)$ -distributed to begin with.

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<sup>2</sup>Recall that the  $\mathcal{N}(\mu, \sigma^2)$  distribution has pdf given by  $f_\mu(x) = \exp(-(x - \mu)^2/2\sigma^2) / \sqrt{2\pi\sigma^2}$ , where  $x \in \mathbb{R}$ .

4. (5 points) Answer each of the following questions by writing YES or NO (1 point), and justify your answer in *at most* three sentences (1.5 points).

(a) Suppose we calculate a 95%-confidence interval  $(L(\mathbf{X}), U(\mathbf{X}))$  for an unknown parameter  $\theta \in \Theta$ , and we observe  $L(\mathbf{X}) = -10$  and  $U(\mathbf{X}) = 7$ . Is it true that  $\mathbb{P}_\theta(-10 \leq \theta \leq 7) = 0.95$ ?

(b) Let  $T(\mathbf{X})$  be sufficient for  $\theta$  and let  $\lambda(\mathbf{X})$  and  $\lambda^*(T(\mathbf{X}))$  be the LRT statistics based on  $\mathbf{X}$  and  $T(\mathbf{X})$ , respectively. If a size- $\alpha$  LRT rejects  $H_0$  when  $\lambda(\mathbf{X}) \leq 0.261$ , will a test that rejects  $H_0$  when  $\lambda^*(T(\mathbf{X})) \leq 0.261$  also be a size- $\alpha$  LRT?

5. (5 points) Let  $X_1, X_2, \dots, X_n \stackrel{iid}{\sim} \text{Bernoulli}(\theta)$ , where  $\theta \in (0, 1)$ .<sup>3</sup> Suppose we want a level- $\alpha$  test of the hypothesis  $H_0 : \theta \leq \theta_0$  versus  $H_A : \theta > \theta_0$ .

(a) (4 points) Using the Karlin-Rubin theorem, show that for some  $c_0 > 0$ , the test with rejection region  $R = \{\mathbf{x} \in \mathcal{X}^n : \sum_{i=1}^n x_i > c_0\}$  is a UMP level- $\alpha$  test.

(b) (1 point) How might you find an appropriate  $c_0$  for the test above?

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<sup>3</sup>Recall that the Bernoulli( $\theta$ ) distribution has pmf given by  $f_\theta(x) = \theta^x(1-\theta)^{1-x}$  for  $x \in \{0, 1\}$ . It might also help to recall that the Bin( $n, \theta$ ) distribution has pmf given by  $\binom{n}{x}\theta^x(1-\theta)^{n-x}$  for  $x \in \{0, 1, \dots, n\}$ .

6. (5 points) Suppose  $\boldsymbol{\theta} = (\theta_1, \dots, \theta_k) \in \mathbb{R}^k$  and that we can produce a  $(1 - \alpha_j)$ -confidence interval  $(L_j(\mathbf{X}), U_j(\mathbf{X}))$  for each  $\theta_j$ . Prove that if we choose  $\alpha_1 = \dots = \alpha_k = \alpha/k$ , then

$$C(\mathbf{X}) = (L_1(\mathbf{X}), U_1(\mathbf{X})) \times \dots \times (L_k(\mathbf{X}), U_k(\mathbf{X}))$$

is a  $(1 - \alpha)$ -confidence region for  $\boldsymbol{\theta}$ , in the sense that  $\mathbb{P}_{\boldsymbol{\theta}}(\boldsymbol{\theta} \in C(\mathbf{X})) \geq 1 - \alpha$ .

*Hint:* One of De Morgan's laws implies that for any sets  $A_1, \dots, A_n$ , we have  $\cap_{i=1}^n A_i = (\cup_{i=1}^n A_i^c)^c$ .

7. (BONUS: 5 points) Suppose that  $X_1, \dots, X_n \stackrel{iid}{\sim} F_\theta$  are continuous and that  $\theta$  is the *median* of  $F_\theta$ : that is,  $\mathbb{P}_\theta(X_i \leq \theta) = 1/2$ . Find the smallest sample size  $n$  such that  $(X_{(1)}, X_{(n)})$  is a  $(1 - \alpha)$ -confidence interval for  $\theta$ .