STA2311: Advanced Computational Methods for Statistics I Class 4: Stochastic Optimization

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1/41









Section 1

Introduction

Assumed Knowledge

- In this class and the next, assume we know how to sample simple random variables
- Essentially what was discussed in Class 1
- Usually the multivariate normal distribution will suffice

```
▶ To sample \pmb{X} \sim \mathcal{N}_d(\pmb{\mu}, \pmb{\Sigma}) in R, use
```

• We will discuss sampling much more thoroughly in Class 6

4/41

Deterministic vs. Stochastic Algorithms

- In the last two classes, we learned about classical optimization techniques and the EM algorithm
- These algorithms were all deterministic
- That is, each algorithm, run with the same inputs (known parameters, initializations, etc.) would invariably produce the same outputs
- It turns out that introducing randomness into certain optimization procedures can dramatically improve their performance
- Such algorithms are generally called *stochastic optimization procedures*

A Simple Stochastic Algorithm

- ullet To illustrate, suppose we want to minimize a function $g:\mathbb{R}^d\to\mathbb{R}$
- Given a current best guess θ_t of the minimizer, we can propose a candidate $\theta' = \theta_t + Z$, where $Z \sim N_d(\mathbf{0}, \sigma^2 \mathbf{I})$ for some $\sigma^2 > 0$
- If $g(\theta') < g(\theta_t)$, then accept the candidate and set $\theta_{t+1} = \theta'$; otherwise, set $\theta_{t+1} = \theta_t$
- This is called random optimization, and works well when d is small
 - But performance becomes much worse as d increases due to the curse of dimensionality
- We will discuss several more refined stochastic optimization methods

Section 2

Stochastic Gradient Descent

Gradient Descent

- Recall the gradient descent (GD) algorithm from Class 2
- ullet We aim to minimize a differentiable function $g:\mathbb{R}^d\mapsto\mathbb{R}$
- Starting with an initial value $heta_0$, the basic GD procedure selects

$$\boldsymbol{\theta}_{t+1} = \boldsymbol{\theta}_t - h \nabla g(\boldsymbol{\theta}_t)$$

for some pre-chosen h > 0

• Once a stopping criterion has been met (say at iteration T), we could output θ_T or $\frac{1}{T} \sum_{t=1}^{T} \theta_t$ or θ_s where $s = \underset{t \leq T}{\operatorname{argmin}} g(\theta_t)$

8/41

The Problem with GD

- In statistical applications we are interested in minimizing $g(\theta) = -l(\theta|\mathbf{y}) = -\sum_{i=1}^{n} \log(f(y_i|\theta))$ which we often do by solving $0 = \nabla g(\theta) = -\frac{1}{n} \sum_{i=1}^{n} \frac{\partial}{\partial \theta} \log(f(y_i|\theta))$
- For example, in multiple linear regression, the function to be minimized is

$$g(\boldsymbol{\theta}) = \frac{1}{n} \sum_{i=1}^{n} (y_i - \boldsymbol{\theta}^\top \boldsymbol{x}_i)^2$$

and its gradient is

$$\nabla g(\boldsymbol{\theta}) = \frac{2}{n} \sum_{i=1}^{n} \left((\boldsymbol{\theta}^{\top} \boldsymbol{x}_{i} - y_{i}), x_{i,1} (\boldsymbol{\theta}^{\top} \boldsymbol{x}_{i} - y_{i}), \dots, x_{i,p} (\boldsymbol{\theta}^{\top} \boldsymbol{x}_{i} - y_{i}) \right)$$
(1)

where n is the number of observations in our training dataset

From GD to SGD

- If sample size n is large, the computational cost is too large so instead one could use a random sample from {y₁, y₂,..., y_n} to produce an unbiased estimator of ∇g(θ)
- In stochastic gradient descent (SGD), we replace the gradient $\nabla g(\theta^{(t)})$ with a random vector $W^{(t)}$
- The random vector is chosen so that its expected value is $abla g(m{ heta}^{(t)})$
- In other words, we seek an unbiased estimator of the gradient
- Canonical choices:
 - $\boldsymbol{W}^{(t)} = \nabla_{\theta} \log(f(y_{I}|\theta))$ where $I \sim Uniform\{1, \ldots, n\}$
 - ► $\boldsymbol{W}^{(t)} = \frac{1}{K} \sum_{j=1}^{K} \nabla_{\theta} \log(f(y_{i_j}|\theta))$ where $\{i_1, \ldots, i_K\}$ is a simple random sample (without replacement) from $\{1, \ldots, n\}$.

Loss Functions

• In general, suppose that the function to be minimized is

$$g(\boldsymbol{\theta}) = \frac{1}{n} \sum_{i=1}^{n} L(\boldsymbol{\theta}, \boldsymbol{y}_i),$$

where y_i is the *i*'th observation in our dataset and *L* is a *loss function* • Therefore

$$\nabla g(\boldsymbol{\theta}) = \frac{1}{n} \sum_{i=1}^{n} \nabla_{\boldsymbol{\theta}} L(\boldsymbol{\theta}, \boldsymbol{y}_i) = \mathbb{E}_I \left[\nabla_{\boldsymbol{\theta}} L(\boldsymbol{\theta}^{(t)}, \boldsymbol{Y}_I) \right]$$

• The random vector $\boldsymbol{W}^{(t)}$ we choose satisfies

$$\mathbb{E}\left[\boldsymbol{W}^{(t)} \mid \boldsymbol{\theta}^{(t)}\right] = \mathbb{E}_{I}\left[\nabla_{\boldsymbol{\theta}} L(\boldsymbol{\theta}^{(t)}, \boldsymbol{Y}_{I})\right]$$

where the left side expectation is with respect to whatever randomization procedure we have to determine \boldsymbol{W} .

11/41

Unbiased Estimators of the Gradient

- Use a single-observation-at-a-time design and cycle through all observations
- Partition data into batches of size *K* << *n* which are sampled at random without replacement.

The Algorithm

- The stochastic gradient descent algorithm is
 - Initialize the process at θ_0 and choose a pre-specified step size h > 0and number of iterations T
 - So Make the updates $\theta_{t+1} = \theta_t h_t W^{(t)}$ for $0 \le t \le T$, where $W^{(t)}$ is a random vector such that $\mathbb{E} \left[W^{(t)} \mid \theta^{(t)} \right] = \nabla g(\theta^{(t)})$
 - Stepsize h_t can be constant or can be decreased $h_t = \gamma^t h$ with $\gamma \in (0, 1)$.
- Clearly, the update is no longer guaranteed to decrease g at every update.
- Convergence is towards a ball centered at the stationary point θ^* so the method is often used when approximate rather than precise optimization is acceptable (e.g., ML applications with massive volume of training data)
- Theory...

Example: Logistic Regression

• Consider the same logistic regression example from Class 2

```
set.seed(7)
expit <- function(x) \{1/(1+exp(-x))\}
logit \leftarrow function(p) \{ log(p/(1-p)) \}
norm <- function(x) {sqrt(sum(x^2))}</pre>
n <- 1000
X1 \leq rnorm(n=n)
X2 <- rbinom(n=n, size=1, prob=0.2)
X3 <- rpois(n=n, lambda=0.7)
X \leftarrow cbind(1, X1, X2, X3)
y <- rbinom(n=n, size=1, prob=expit(0.4 + 0.7*X1 + 3*X2 - X3))</pre>
grad.g <- function(theta, XX, yy) {
  t(XX) %*% (expit(apply(XX, 1, function(x) x%*%theta)) - yy)}
```

Example: Logistic Regression (Continued)

```
# use TT iterations
TT <- 1000
# size of minibatch
k <- 50</pre>
```

```
#want to go through all the data in rep1 updates
repeats1=n/k
#need to cycle through all the data rep2 times
repeats2=TT/repeats1
```

```
sub.sample=matrix(0,nrow=TT, ncol=k)
for(i in 1:repeats2){
    samp.inds<-sample(1:n)
sub.sample[((i-1)*repeats1+1):(i*repeats1),]=
    matrix(samp.inds,ncol=k,nrow=repeats1,byrow=T) }</pre>
```

```
#stores the parameter values
ths <- cbind( rep(1, 4), matrix(0L, nrow=4, ncol=TT-1))</pre>
```

15/41

Example: Logistic Regression (Continued)

```
for (t in 1:(TT-1)) {
 # samp.inds <- sample(1:n, size=k,replace=F)</pre>
  X.k <- X[as.vector(sub.sample[t,]),]</pre>
  y.k <- y[as.vector(sub.sample[t,])]</pre>
   #fixed stepsize
  alp <- 0.005
  ths[,t+1] \leftarrow ths[,t] - alp*grad.g(ths[,t], X.k, y.k)
}
th.SGD <- ths[,TT]
th.NR <- as.vector(glm(y ~ ., family = binomial(link="logit"),
                         data=data.frame(y, X1, X2, X3))$coefficients)
print(cbind(th.SGD,th.NR))
par(mfrow=c(2,2))
for (i in 1:4) ts.plot(ths[i,])
```

Section 3

Simulated Annealing

Basic Metallurgy

- Reference: Goffe et al. [1994]
- In metallurgy (or thermodynamics more generally), *annealing* means "slow cooling"
- When casting metallic objects from molten metal, the final goal is to bring the metal to a minimum-energy state (where it is very hard)
- However, hot metal is easier to mold
- One wants to shape the metal while slowly cooling it (cooling it too fast will not allow reaching the desired shape)
- ... but not too slowly, since we don't want to wait too long
- Simulated annealing seeks to minimize/maximize a function according to the same principle

The set up

- Interested in minimizing $g(\theta) : E \subset \mathbb{R}^d \to \mathbb{R}$ (or maximizing $-g(\theta)$) where g is not restricted (can be discrete, non-continuous, etc).
- We will assume here that g is bounded
- Note that max of -g is also max of $\tilde{g} = \exp(-g)$ which can be thought of as an unnormalized density on E.
- Most of the algorithms studied so far have trouble escaping a local extrema point
- We "heat" $\tilde{g} = \exp(-g(\theta))$ into $\tilde{g}_t = \exp(-g(\theta)/t)$ where $t \ge 1$.
- Intuition: As $T \to \infty \tilde{g}_T(\theta) = 1_E(\theta)$ so if \tilde{g} has extreme points separated by "big dips" its heated versions are more "level"

19/41

Example

```
f.T=function(nn=3,xx=c(1,1), TT=1){
    rez=(16*xx[1]*(1-xx[1])*xx[2]*(1-xx[2])*sin(nn*pi*xx[1])*sin(nn*pi*xx[2]
    return(exp(rez/TT))}

n.eval=100
co1=ppoints(n.eval)
co3=ppoints(n.eval)
TT=1
for(i in 1:n.eval)
co3[i]=f.T(n,c(co1[i],0.5),TT)
plot(co1,co3,col="black",type="l", xlab="x",ylab="f(x,0.5)")
```

Example

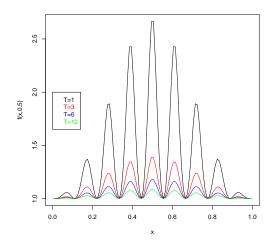


Figure 1: Illustration of annealing

21 / 41

Building Up the Algorithm

- Suppose at time t, our best guess of the optimum is $\theta^{(t)}$
- We construct a candidate θ' for $\theta^{(t+1)}$ by randomly perturbing one element of $\theta^{(t)}$
- If $g(\theta') < g(\theta_t)$, good we've improved upon the old guess, so take $\theta_{t+1} = \theta'$
- Otherwise, θ_t could just be a *local* minimum that we'd like to escape from
- In order to escape, one must accept moves that seem counterproductive (e.g., increase the value of g occasionally)

The Algorithm

- The basic simulated annealing algorithm is
 - Choose a run length M, a cooling schedule $T : \{0, 1, ..., M\} \to (0, \infty)$ and initialize the process at θ_0
 - 2 Evaluate $g(\theta_0)$
 - So For 1 ≤ t ≤ M, propose a new θ' by perturbing a random coordinate: $θ'_j = θ_j + ξ_t \text{ for some random variable } ξ_t$

★ If
$$g(heta') < g(heta_t)$$
, take $heta_{t+1} = heta$

★ If $g(\theta') \ge g(\theta_t)$, set $\theta_{t+1} = \theta_t$ with probability

$$e^{(g(\theta_t)-g(\theta'))/T(t)} = \frac{\tilde{g}_{T(t)}(\theta')}{\tilde{g}_{T(t)}(\theta_t)}$$

and $\theta_{t+1} = \theta_t$ otherwise

• Note that in our example the ratio $\frac{\tilde{g}_{\mathcal{T}(t)}(\theta')}{\tilde{g}_{\mathcal{T}(t)}(\theta_t)}$ will tend to be closer to 1 when $\mathcal{T}(t) > 1$.

That Random Variable

- The random variable ξ_t used to perturb the coordinate at step t of the algorithm is chosen based on what we know about the objective function
- If g is continuous with domain ℝ^d, then we can take ξ_t ^{iid} N(0, σ²) for some chosen variance σ²
- Or the the algorithm can be made to be adaptive, with independent $\xi_t \sim \mathcal{N}(0, \sigma_t^2)$
- But the domain of g can be finite (like the configuration of atoms in a metal) too
 - In this case, the "perturbed' coordinate can be chosen randomly from the configuration space of g

The Cooling Schedule

- The "temperature'' evolves according to the cooling schedule (or annealing schedule) and should generally decrease to 0 as $t \to \infty$
- There are many choices for the functional form of T(t)
- Linear: T(t) = T(0) dt for some d > T(0)/M
- Logarithmic: $T(t) = T(0)/\log(1+t)$
- Geometric: $T(t) = r \cdot T(t-1)$ for some $r \in (0,1)$
- Exponential: $T(t+1) = (T(1)/T(0))^t \cdot T(t)$

Example: Logistic Regression (Again)

```
set.seed(2311)
expit <- function(x) \{1/(1+exp(-x))\}
logit \leftarrow function(p) \{ log(p/(1-p)) \}
norm <- function(x) {sqrt(sum(x^2))}
n <- 1000
X1 \leq rnorm(n=n)
X2 <- rbinom(n=n, size=1, prob=0.2)
X3 <- rpois(n=n, lambda=0.7)
X \leftarrow cbind(1, X1, X2, X3)
y <- rbinom(n=n, size=1, prob=expit(0.4 + 0.7*X1 + 3*X2 - X3))</pre>
g <- function(theta) {
  -sum(y*log(expit(apply(X, 1, function(x) x%*%theta))) +
       (1-y)*log(1-expit(apply(X, 1, function(x) x%*%theta))))}
```

```
Example: Logistic Regression (Again) (Continued)
M <- 5000
th <- rep(1, times=4)
for (t, in 1:M) {
  TT <- 10/log(t+1)
  gth <- g(th)
  p.ind <- sample(1:4, size=1)</pre>
  th.p <- th
  th.p[p.ind] \leftarrow th.p[p.ind] + rnorm(n=1, sd=.5)
  gth.p <- g(th.p)
  if (gth.p < gth || runif(n=1) < min(exp(gth - gth.p)/TT,1)) {
    th <- th.p
  } else {
    th <- th
  }
}
th.SA <- th
```

Benefits

- Observe that there are essentially no restrictions on the objective function g
- It need not be continuous, which makes simulated annealing useful for discrete optimization problems
 - The *travelling salesman problem* is a classical example
- We will see later in the course that simulated annealing is a particular kind of Markov Chain Monte Carlo (MCMC) algorithm

Section 4

Genetic Algorithms

Inspired by Evolution

- Genetic Algorithms (GA) are iterative stochastic algorithms often used in discrete optimization
- The design is inspired by theory of evolution and adaptation through genetics
- Assumptions:
 - ► (A1) Fitness: the quality of a potential solution can be evaluated using a fitness function, e.g. if the problem of interes is finding arg max_x g(x) then the fitness of a candidate solution x₀ is g(x₀).
 - (A2) Representation: every candidate solution v to the optimization problem can be represented as a string of bits (vectors whose entries are 0 or 1) v ∈ {0,1}^M

Implementation

• The algorithm cycles through the following steps:

- From a population of candidate solutions at iteration t, S_t = {v_i, g(v_i) : 1 ≤ i ≤ K} we use selection to produce an intermediate population
- The intermediate population can be created in different ways:
 - \star It is S_t
 - * Via sampling K times with replacement from S_t using $p_i \propto g(v_i)$ as the sampling weights.
 - ***** Use the fittest *L* elements in S_t and K L random elements from S_t .
- ► The population at time *t* + 1 is obtained from the intermediate population using crossover and mutation.

Crossover

• Consider two potential solutions/strings

```
(101...1101) and (yxy...xxy)
```

- Randomly select the crossover point $c \in \{1,\ldots, {\cal K}-1\}$
 - 11010 \/ 01100101101
 - yxyyx / yxxyyxyxy
- 11010yxxyyyxyxy and yxyyx01100101101
- There can be more than one crossover point

Mutation

- \bullet With certain probability ϵ a random element in the string is flipped
- Alternatively, can think of mutation of generating a random new bit for a randomly selected component of *v_i*
- In summary: for each *i* flip a coin to decide whether *v_i* will mutate. If it does select the component uniformly at random and with probability *ε* flip it.

Possible references: Whitley [1994] Chatterjee et al. [1996]

Example

```
• Find arg max<sub>(x_1,x_2)</sub> f(x_1,x_2) where
            f(x_1, x_2) = [16x_1x_2(1 - x_1)(1 - x_2)\sin(9\pi x_1)\sin(9\pi x_2)]^2
library(rgl)
f=function(nn=3,xx=c(1,1))
    rez=(16*xx[1]*(1-xx[1])*xx[2]*(1-xx[2])*
            sin(nn*pi*xx[1])*sin(nn*pi*xx[2]))^2
    return(rez)}
# we want to plot f
n.eval=100
co1=ppoints(n.eval)
co2=ppoints(n.eval)
s=matrix(0,n.eval,n.eval)
n=9
set.seed(17)
```

```
for(i in 1:n.eval){for(j in 1:n.eval)
  {s[i,j]=f(n,c(co1[i],co2[j]))}}
persp3d(co1,co2,s,col = "white", package = "rgl")
contour(co1,co2,s)
S=80 # size of each generation
itr= 100 # number of iterations to run
m=10 # the number of digits (determines the precision of solution)
m.rate=0.01 # mutation rate
rnk = c(1:S) # ranks the specimens in one generation
phi =c(1:S) # fitness for each specimen
best.ftness =c(1:(itr+1)) # best fitness found
ftness =matrix(0,nrow=(itr+1),ncol=S) # stores all the fitness-es
bin.current=array(0,c(S,2,m)) # current population
bin.next=array(0,c(S,2,m)) # current population
twos=(1/2)^{c}(1:m)
x.current=x.next=matrix(0,ncol=2,nrow=S)
```

```
# initialize the population
```

```
for(i in 1:S){
bin.current[i,1,]=rbinom(m,1,0.5)
bin.current[i,2,]=rbinom(m,1,0.5)
x.current[i,1]=sum(twos*bin.current[i,1,])
x.current[i,2]=sum(twos*bin.current[i,2,])
phi[i]=f(n,c(x.current[i,1],x.current[i,2]))}
```

```
ftness[1,]=phi
rnk=order(phi)
best.ftness[1]=max(phi)
```

```
Example (cont'd)
for(j in 1:itr-1)
{
```

BUILDS THE NEW GENERATION, SELECTING FIRST PARENT BASED ON
 # FITNESS AND THE SECOND PARENT AT RANDOM
 # THERE ARE S/2-2 BREEDINGS ALLOWED RESULTING IN S OFFSPRINGS

```
# we keep the top two specimens in the population
 for(i in 1:S){
 x.current[i,1]=sum(twos*bin.current[i,1,])
 x.current[i,2]=sum(twos*bin.current[i,2,])
 phi[i]=f(n,c(x.current[i,1],x.current[i,2]))
 }
 ftness[1,]=phi
 rnk=order(phi)
 bin.next[1,,]=bin.current[rnk[S],,]# this one has the highest rank in fitne
 bin.next[S,,]=bin.current[rnk[S-1],,] # has the second highest rank in fitr
 x.current[1,]=x.current[rnk[S],]
Radu Craiu, Robert Zimmerman (UofT)
                            STA2311: Computational Methods I
                                                             October 3, 2023
                                                                             37 / 41
```

```
for(i in 2:(S/2)) {
# samples from {1,...,S} with probs prop to phi
parent1.index=sample(1:S,1,prob=phi)
parent2.index = sample(1:S,1)
#crossover position
pos = sample(1:(m-1), 1)
#the first offspring is produced via crossover
# first coordinate
bin.next[i,1,1:pos]=bin.current[parent1.index,1,1:pos]
bin.next[i,1,(pos+1):m]=bin.current[parent2.index,1,(pos+1):m]
#second coordinate
bin.next[i,2,1:pos]=bin.current[parent1.index,2,1:pos]
bin.next[i,2,(pos+1):m]=bin.current[parent2.index,2,(pos+1):m]
```

```
# THE MUTATION STEP IS PERFORMED
mutate = rbinom(m,1,m.rate)
#if a mutation has occured, the coordinate is flipped
bin.next[i,1,] = (bin.next[i,1,]+mutate)%/2
# repeat the process for second offspring
pos = sample(1:(m-1), 1)
mutate = rbinom(m,1,m.rate)
bin.next[S-i+1,1,1:pos]=bin.current[parent2.index,1,1:pos]
bin.next[S-i+1,1,(pos+1):m]=bin.current[parent1.index,1,(pos+1):m]
bin.next[S-i+1,2,1:pos]=bin.current[parent2.index,2,1:pos]
bin.next[S-i+1,2,(pos+1):m]=bin.current[parent1.index,2,(pos+1):m]
bin.next[i,1,] = (bin.next[i,1,]+mutate)%2
for(i in 2:(S-1)){
x.current[i,1]=sum(twos*bin.next[i,1,])
x.current[i,2]=sum(twos*bin.next[i,2,])
phi[i]=f(n,c(x.current[i,1],x.current[i,2]))
}
bin.current=bin.next
```

```
# update the fitness values
for(k in 1:S){
    x.current[k,1]=sum(twos*bin.current[k,1,])
    x.current[k,2]=sum(twos*bin.current[k,2,])
   phi[k]=f(n,c(x.current[k,1],x.current[k,2]))
                                                    }
ftness[j+1,]=phi
best.ftness[j+1]=max(phi)}
print(x.current[rnk[S],])
plot(c(1,40),c(0,1.1),type="n", xlab="Generation", ylab="Best solution")
lines(c(1:itr),best.ftness[1:itr])
abline(h=f(n,c(0.5,0.5)),col="red")
```

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- Darrell Whitley. A genetic algorithm tutorial. *Statistics and computing*, 4: 65–85, 1994.